Markov Models

Lecture 10


## The honest and the dishonest casino



## We can use the conditional probabilities for discrimination



We can just compare $P(M$ and model $L)$ and $P(M$ and model $F)$

## We can use the conditional probabilities for discrimination


$\mathrm{P}(\mathrm{M}$ and model L$)=0.5^{*} 0.5^{*} 0.5^{*} 0.1^{*} 0.5^{*} 0.1=0.000625$
$P(M$ and model $F)=0.17^{*} 0.17^{*} 0.17^{*} 0.17^{*} 0.17^{*} 0.17=0.000024$

|  | F | L |
| :--- | :--- | :--- |
| 1 | 0.17 | 0.10 |
| 2 | 0.17 | 0.10 |
| 3 | 0.17 | 0.10 |
| 4 | 0.17 | 0.10 |
| 5 | 0.17 | 0.10 |
| 6 | 0.17 | 0.50 |

How confident we are that this sequence was produced by a loaded die?
$P(M$ and model L$) / \mathrm{P}(\mathrm{M}$ and model F$)=25.89$
Or $\log [P(M$ and model $L) / P(M$ and model $F)]=1.4$

## The occasionally dishonest casino



# Sequence generated by a model of an occasionally dishonest casino 



## Markov chains

- A general model of a system which moves from state to state with some probability $a_{i j}$, called a transition probability
- While in a particular state, system emits a symbol $m_{k}$ from a finite alphabet with the probability $e_{i}\left(m_{k}\right)$, called an emission probability of symbol $m_{k}$ in state Wi
- If we construct the schedule of observation times and at each point in time record the symbols emitted by a system along with the state, we obtain 2 sequences: the sequence of emitted symbols which is called an observed sequence M , and the sequence of states which is called a path through system states


## Markov chain terminology

Transition probabilities


## Markov chain terminology

Emission probabilities


## Markov model diagram



## Markov model parameters

Emission probabilities
The transition matrix

|  | $F$ | $L$ |
| :--- | :--- | :--- |
| $F$ | 0.83 | 0.17 |
| $L$ | 0.60 | 0.40 |


|  | F | L |
| :--- | :--- | :--- |
| 1 | 0.17 | 0.10 |
| 2 | 0.17 | 0.10 |
| 3 | 0.17 | 0.10 |
| 4 | 0.17 | 0.10 |
| 5 | 0.17 | 0.10 |
| 6 | 0.17 | 0.50 |

## Hidden Markov Model (HMM)



- States are unknown (hidden)


## Questions to HMM

- Given a sequence of observations, what is the most probable sequence of the underlying states (Most probable path)
- Given a sequence of N observations, what is the probability of obtaining this sequence given a model described by a particular HMM (Sequence probability)
- Given a sequence of N observations, what is the probability that the i-th observation was produced when the system was in state Wj


## The probability that the sequence was generated given a particular path

- Pick the path $\pi$
- The probability $\mathrm{P}(\mathrm{M} \mid \pi)$ is the conditional probability that sequence M was generated while system was moving from state to state according to $\pi$


## The probability that the sequence $M$ was generated following a path $\pi$

- Pick a path $\pi$
- Calculate a joint probability of $\pi$ and $M$

A suggested path

$P(M$ and $\pi)=0.17{ }^{*} 0.83^{*} 0.17$ * $0.17^{*} 0.50$ * $0.60 * 0.50=0.0006$

|  | $F$ | $L$ |
| :--- | :--- | :--- |
| $F$ | 0.83 | 0.17 |
| $L$ | 0.60 | 0.40 |

- Repeat for each possible path and choose a path which maximizes
$P(\pi$ and $M)$. Total $2^{N}$ calculations


## Viterbi algorithm for the most probable path

Dynamic programming

## Dynamic programming. Initialization the probability of choosing a die for the first time

- Add to the system a start state and parameters - the probabilities of choosing a fair or a loaded die in the beginning of a game


State F (fair die)
State L (loaded die)

# Dynamic programming. Initialization 

The graph of a process.


## Dynamic programming. Recursion

The graph of a process. We are looking for a path which maximizes the probability of emission M


# Dynamic programming. Recursion 

If we know the best paths ending at states $L$ and $F$ in position 4 , we can choose max between them and terminate the program


## Dynamic programming. Recursion

This can be repeated for each combination of a position in a sequence of observations and one of 2 states


Note: the probabilities are multiplied, not added up

## Viterbi algorithm. Demo 1



We have reached position $i=1$ with the probability $0.9^{*} 0.17$ of

|  | $F^{\prime}$ | $L$ |
| :--- | :--- | :--- |
| 1 | 0.17 | 0.10 |
| 2 | 0.17 | 0.10 |
| 3 | 0.17 | 0.10 |
| 4 | 0.17 | 0.10 |
| 5 | 0.17 | 0.10 |
| 6 | 0.17 | 0.50 |
|  | $F$ | $L$ |
| F | 0.83 | 0.17 |
| $L$ | 0.60 | 0.40 |
| 0 | 0.90 | 0.10 | going to the F state and emitting 3 , and with probability $0.1^{*} 0.10$ of going to the L-state and emitting 3 . There are no other possibilities

## Viterbi algorithm. Demo 2

 between these two: $0.15^{*} 0.83^{*} 0.17=0.002$

The L-state in position $\mathrm{i}=2$ can be reached with probability $0.01^{*} 0.40^{*} 0.10$ or $0.15^{*} 0.17^{*} 0.10=0.0026$. The second is larger so we choose it.

## Viterbi algorithm. Demo 3



We can reach position $\mathrm{i}=3$ ( F -state) with the probability $0.02 * 0.83^{*} 0.17=0.0028$ or with probability

|  | $F^{\prime}$ | $L$ |
| :--- | :--- | :--- |
| 1 | 0.17 | 0.10 |
| 2 | 0.17 | 0.10 |
| 3 | 0.17 | 0.10 |
| 4 | 0.17 | 0.10 |
| 5 | 0.17 | 0.10 |
| 6 | 0.17 | 0.50 |
|  | $F$ | $L$ |
| F | 0.83 | 0.17 |
| L | 0.60 | 0.40 |
| 0 | 0.90 | 0.10 | $0.0026^{*} 0.4^{*} 0.17=0.00018$. We chose the max between these two: $0.02^{*} 0.83^{*} 0.17=0.0028$

The L-state in position $i=3$ can be reached with probability $0.02^{*} 0.17^{*} 0.50=0.0017$ or $0.0026^{*} 0.4^{*} 0.5=0.0017$. We chose the second - arbitrarily

## Viterbi algorithm. Demo 4

We can reach position $\mathrm{i}=4$ ( F -state) with the probability $0.0028 * 0.83 * 0.17=0.0004$ or with probability

|  | $F^{\prime}$ | L |
| :--- | :--- | :--- |
| 1 | 0.17 | 0.10 |
| 2 | 0.17 | 0.10 |
| 3 | 0.17 | 0.10 |
| 4 | 0.17 | 0.10 |
| 5 | 0.17 | 0.10 |
| 6 | 0.17 | 0.50 |
|  | F | L |
| F | 0.83 | 0.17 |
| L | 0.60 | 0.40 |
| 0 | 0.90 | 0.10 | $0.0017 * 0.6^{*} 0.17=0.00017$. We chose the max between these two: $0.0028 * 0.83 * 0.17=0.0004$

The L-state in position $\mathrm{i}=4$ can be reached with probability $0.0017 * 0.40 * 0.50=0.00034$ or $0.0028 * 0.17 * 0.5=0.00024$. We chose the max: $0.0017 * 0.40 * 0.50=0.00034$

## Viterbi algorithm. Demo - end



Choose max: 0.0004 . So, the most probable sequence of states: FFFF

|  | $F^{\prime}$ | $L$ |
| :--- | :--- | :--- |
| 1 | 0.17 | 0.10 |
| 2 | 0.17 | 0.10 |
| 3 | 0.17 | 0.10 |
| 4 | 0.17 | 0.10 |
| 5 | 0.17 | 0.10 |
| 6 | 0.17 | 0.50 |
|  | $F$ | $L$ |
| F | 0.83 | 0.17 |
| L | 0.60 | 0.40 |
| 0 | 0.90 | 0.10 |

Evidently, it is not enough to have 2 sixes in a row in order to be able to spot the loaded die.

## Viterbi algorithm. Log-values

$$
\begin{aligned}
& P\left(\pi_{F, 1}\right)=a_{0 F}{ }^{*} e_{F}(M[1]) \quad P\left(\pi_{L, 1}\right)=a_{0 L}{ }^{*} e_{L}(M[1]) \\
& P\left(\pi_{F, i+1}\right)=\max \left\{P\left(\pi_{F, i}\right)^{*} a_{F F}, P\left(\pi_{L, 1}\right)^{*} a_{L F}\right\}^{*} e_{F}(M[i+1]) \\
& P\left(\pi_{L, i+1}\right)=\max \left\{P\left(\pi_{L, i}\right)^{*} a_{L L}, P\left(\pi_{F, i}\right)^{*} a_{F L}\right\}^{*} e_{L}(M[i+1]) \\
& P\left(\pi^{*}\right)=\max \left\{P\left(\pi_{F, N}\right), P\left(\pi_{L, N}\right)\right\}
\end{aligned}
$$

In order to avoid the underflow errors, in practice log is used instead of the actual probabilities

$$
\begin{aligned}
& P\left(\pi_{F, 1}\right)=\log a_{0 F}+\log e_{F}(M[1]) \quad P\left(\pi_{L, 1}\right)=\log a_{0 L}+\log e_{L}(M[1]) \\
& P\left(\pi_{F, i+1}\right)=\max \left\{P\left(\pi_{F, i}\right)+\log a_{F F}, P\left(\pi_{L, 1}\right)+\log a_{L F}\right\}+\log e_{F}(M[i+1]) \\
& P\left(\pi_{L, i+1}\right)=\max \left\{P\left(\pi_{L, i}\right)+\log a_{L L}, P\left(\pi_{F, i}\right)+\log a_{F L}\right\}+\log e_{L}(M[i+1]) \\
& P\left(\pi^{*}\right)=\max \left\{P\left(\pi_{F, N}\right), P\left(\pi_{L, N}\right)\right\}
\end{aligned}
$$

## Viterbi algorithm. Log-values

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& P\left(\pi_{F, i+1}\right)=\max \left\{P\left(\pi_{F, i}\right)^{*} a_{F F}, \quad P\left(\pi_{L, 1}\right)^{*} a_{L F}\right\}^{*} e_{F}(M[i+1]) \\
& P\left(\pi_{L, i+1}\right)=\max \left\{P\left(\pi_{L, i}\right)^{*} a_{L L}, P\left(\pi_{F, i}\right)^{*} a_{F L}\right\}^{*} e_{L}(M[i+1]) \\
& P\left(\pi^{*}\right)=\max \left\{P\left(\pi_{F, N}\right), P\left(\pi_{L, N}\right)\right\}
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& P\left(\pi_{F, i+1}\right)=\max \left\{P\left(\pi_{F, i}\right)+\log a_{F F}, P\left(\pi_{L, 1}\right)+\log a_{L F}\right\}+\log e_{F}(M[i+1]) \\
& P\left(\pi_{L, i+1}\right)=\max \left\{P\left(\pi_{L, i}\right)+\log a_{L L}, P\left(\pi_{F, i}\right)+\log a_{F L}\right\}+\log e_{L}(M[i+1]) \\
& P\left(\pi^{*}\right)=\max \left\{P\left(\pi_{F, N}\right), P\left(\pi_{L, N}\right)\right\}
\end{aligned}
$$

## How good is the prediction

| Rolls | 31511524644654424531132163116415213352514454 631655526566666 |  |
| :---: | :---: | :---: |
| Die |  |  |
| Viterbi | FFF FFFFPFFPFPFFFFFFFFFFFFFFFFFFFFFFFFFEFFFFIFFFLLLLLLLLLLLL | delay |
| Rolls | 651166453132551245636564631636663 /62326455236266666525151631 |  |
| Die |  |  |
| Viterbi | LLLLLLFFFFFFFFFFFFLLLLLLLLLLLLLLLLALLLLLILLLLLLLLLLLLEFFFFFPFF |  |
| Rolls | 222555441656565563564324364131513465.46353411126414626253356 | Missing |
| Die |  | short |
| Viterbi |  | stretches |
| Rolls | 366163666466232534413661661163252562462255265252266435353336 |  |
| Die |  |  |
| viterbi | LLLLLLLLLLLLLFFFFFEFFPFFEFFFFFFEFEYFFFEFEFFFEFEFEFFFEFFFFFFEF |  |
| Ro_ls | 233121625364414432335163243633655624666626326566.2355245242 |  |
| Eie | FFFFPFFFFFFFFPFFFFFFFFFPFFFLLLLLLLLLLLLLLLLLLLLLLFFFFFFFFFFF |  |
| viterbi |  |  |

Overall, an underlying hidden pathway explains the given sequence well - the model is good

## Exercise 1. Markov models

- In Vancouver, if it rains today, then it rains tomorrow 3 times out of 5 . If it is sunny today, it is also sunny tomorrow 1 time out of 3 . Build a Markov model for the weather in Vancouver.


## Exercise 2. Discrimination by probability

- Markov models for the honest and for the dishonest casino are presented below:

$$
\begin{aligned}
& e(\text { Heads })=1 / 2 \\
& e(\text { Tails })=1 / 2
\end{aligned}
$$

Fair coin

$$
\begin{aligned}
& e(\text { Heads })=3 / 4 \\
& e(\text { Tails })=1 / 4
\end{aligned}
$$

Biased coin

Find out what of the coins has more probably produced the following sequence of observations

## Exercise 2. When the coin is

 biased- For sequence M of length N with k heads:
- $P(M$ I fair coin $)=\Pi_{n}(1 / 2)=1 / 2^{N}$
- $P(M \mid$ biased coin $)=\Pi_{k}(3 / 4) * \Pi_{N-k}(1 / 4)=3^{k} / 4^{k *} 1 / 4^{N-k}$
- For this simple model, we can find when
$P(M$ I fair coin $)<P(M \mid$ biased coin)
$1 / 2^{\mathrm{N}}<3^{\mathrm{K}} / 4^{\mathrm{N}}$
$2^{\mathrm{N}}<3 \mathrm{~K}$
$\mathrm{Nlog}_{2} 2<\mathrm{klog}_{2} 3$
$K>\left(\log _{2} 2 / \log _{2} 3\right) N$
$\mathrm{K}>0.63 \mathrm{~N}$


## Exercise 3.

- Using the Viterbi algorithm, find the most probable path of states for the following sequence given the HMM which produced this sequence.


Observed sequence: HTTHHH

## We can answer 2 questions

- What is the probability that a given sequence of observations came from a particular Markov model
- Where in the sequence the model has probably changed


## CpG islands

- C nucleotide followed by G is easily methylated
- Methylated C easily becomes T
- The methylation is suppressed in important regulatory regions - around promoters (starting sites of transcription)
- Thus, an overall low frequency of CG dinucleotide is significantly increased in the gene promoter regions


## Biological questions

- Given a short stretch of DNA sequence, how can we determine whether it came from a CpG island or not
- Given a lon un-annotated DNA sequence, find CpG islands in it


## Markov model for DNA sequence



- Usually, the end of sequence is not modelled in Markov chain - sequence can end anywhere


## Transition probability estimation from real DNA sequences

- From 48 CpG islands of a total length 60,000 nucleotides, and from a regular DNA stretches, the transition probabilities for each pair of nucleotides were estimated (expected 0.25 if at random)
$\mathrm{a}_{\text {from,to }}=$ count $_{\text {from,to }} / \Sigma_{x}$ count $_{\text {from }, x}$

| + | A | C | G | T |
| :--- | :--- | :--- | :--- | :--- |
| A | 0.18 | 0.27 | 0.43 | 0.12 |
| C | 0.17 | 0.37 | 0.27 | 0.19 |
| G | 0.16 | 0.34 | 0.38 | 0.12 |
| T | 0.08 | 0.36 | 0.38 | 0.18 |


| - | $A$ | $C$ | $G$ | T |
| :--- | :--- | :--- | :--- | :--- |
| A | 0.30 | 0.20 | 0.29 | 0.21 |
| C | 0.32 | 0.30 | 0.08 | 0.30 |
| G | 0.25 | 0.25 | 0.30 | 0.20 |
| T | 0.18 | 0.24 | 0.29 | 0.29 |

## Am I in the CpG island?

- To use these (+) and (-) models for discrimination for a given sequence we calculate the log-odds ratio:


## Score(M)=log [ $\mathbf{P ( M | g i v e n ~ m o d e l ~ + ~}) / \mathbf{P}(\mathrm{M} \mid$ given model -)]

If this value is positive, we are in the CpG island, if not, we are not


Test on another set of labeled DNA sequences

## Finding CpG islands - HMM



- The relabeling is the critical step. The essential difference between a Markov chain and an HMM is that for HMM there is no 1-to-1 correspondence between the states and the symbols
- By looking at a single symbol, there is no way to tell whether it came from state C+ or C-


## The most probable path through the sequence of states

- The most probable path for sequence CGCG

| $v$ |  | $C$ | $G$ | $C$ | $G$ |
| :--- | ---: | ---: | ---: | ---: | ---: |
| $\boldsymbol{B}$ | 1 | 0 | 0 | 0 | 0 |
| $\mathbf{A}_{+}$ | 0 | 0 | 0 | 0 | 0 |
| $\mathrm{C}_{+}$ | 0 | 0.13 | 0 | $\mathbf{0 . 0 1 2}$ | 0 |
| $\mathrm{G}_{+}$ | 0 | 0 | $\mathbf{0 . 0 3 4}$ | 0 | $\mathbf{0 . 0 0 3 2}$ |
| $\mathrm{~T}_{+}$ | 0 | 0 | 0 | 0 | 0 |
| $\mathbf{A}_{-}$ | 0 | 0 | 0 | 0 | 0 |
| $\mathrm{C}_{-}$ | 0 | 0.13 | 0 | 0.0026 | 0 |
| $\mathbf{G}_{-}$ | 0 | 0 | 0.010 | 0 | 0.00021 |
| $\mathbf{T}_{-}$ | 0 | 0 | 0 | 0 | 0 |

When we apply the Viterbi algorithm to a long un-annotated DNA sequence, the states will switch between + and - , giving suggested boundaries for CpG islands

## Defining the model for HMM

- 2 parts:
- Model topology: what states there are and how are they connected
- The assignment of parameter values: the transition and emission probabilities


## Parameter estimation

- We are given a set of training sequences
- 2 cases:
- When the states in the training sequences are known
- $\mathrm{a}_{\text {from }, \mathrm{to}}=$ count $_{\text {from }, \mathrm{to}} / \Sigma_{\mathrm{x}}$ count $_{\text {from }, \mathrm{x}}$
- $e_{\text {state } i}($ Symbol $j)=$ count $_{\text {state } i}($ symbol $j) / \Sigma_{y}($ symbol $y)$
- When the states are unknown
- Viterbi training


# Parameter estimation when the states are known - example 

| X | 1 | 2 | 6 | 6 | 1 | 1 | 2 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\pi$ | F | L | F | F | L | L | L |

$$
e_{F}(3)=0 ?
$$

To avoid this, use pseudocounts
$e_{F}(1)=(1+1) /(3+6), 1$ is a pseudocount, 6 is the number of different symbols

$$
e F(1)=2 / 9
$$

$$
e_{F}(2)=1 /(3+6)=1 / 9
$$

$$
e_{F}(3)=1 /(3+6)=1 / 9
$$

$$
e_{F}(4)=1 /(3+6)=1 / 9
$$

$$
e_{F}(5)=1 /(3+6)=1 / 9
$$

$$
e_{F}(6)=(2+1) /(3+6)=3 / 9
$$

$$
\begin{aligned}
& a_{F, L}=2 / 3 \\
& a_{F, F}=1 / 3 \\
& a_{L, F}=1 / 3 \\
& a_{L, L}=2 / 3
\end{aligned}
$$

Or with pseudocounts

$$
\begin{aligned}
& a_{F, L}=2+1 / 3+2=3 / 5 \\
& a_{F, F}=1+1 / 3+2=2 / 5 \\
& a_{\mathrm{L}, \mathrm{~F}}=1+1 / 3+2=2 / 5 \\
& a_{\mathrm{L}, \mathrm{~L}}=2+1 / 3+2=3 / 5
\end{aligned}
$$

## Viterbi training for parameter estimation

- Pick a set of random parameters
- Find the most probable path of states according to this set of parameters
- This path partitions the sequences into partitions according to the states
- Calculate new set of parameters, now from the known states
- Repeat - find the most probable path with the new parameters etc. - until the path does not change anymore


## Viterbi training

- The assignment of paths is a discrete process, thus the algorithm converges precisely.
- When there is no path change, the parameters will not change either, because they are determined completely by the paths
- The algorithm maximizes the probability P (observed data| $\Theta, \pi^{*}$ )
and not $P($ observed data $\mid \Theta)$ which we ideally want


## Parameter estimation illustration 1



The parameters estimated from 300 random rolls and an iterative process started from randomly selected parameters

## Parameter estimation illustration 2



The parameters estimated from 30000 random rolls and an iterative process started from randomly selected parameters

